

Multi-Objective Compromise Allocation Stratified Sampling in the Presence of Non-Response Using Quadratic Cost Function

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Abstract

In this paper we propose a quadratic cost function for allocating sample size in multivariate stratified random sampling in the presence of the non-response. We use the separate linear regression estimator. In this multi-objective non-linear integer programming problem, we use extended lexicographic goal programming for solution purpose. To illustrate the application, we apply this formulation on a real data set.

Keyword: Multivariate stratified sampling, Non-response, Compromise allocation, Multi-objective programming, Quadratic cost function, Lexicographic goal programming

1- Introduction

A good sampling plan plays a significant role to make the results useful, obtained from statistical studies and provides close approximation to the population estimates. A suitably selected sampling plan and samples, representing population, produce more reliable estimates.

The important consideration in stratified random sampling is the sample size allocation in each stratum with the criteria either to minimize variance of stratified sample mean for a fixed cost or to minimize cost for the specified variance.

Stratified random sampling is used to increase precision following some cost mechanism. Allocation of sample size n_h to individual stratum becomes more complicated in a study or survey in practical utilization of stratified random sampling scheme. The researcher meets problem to select a sample that maximizes precision of finite population mean under cost constraint.

Sampling efficiency depends largely on how the sample size is allocated. In univariate stratified sampling, individual optimum allocation can be used when the characteristics are correlated but in case when the characteristics are uncorrelated a suitable criterion is needed for allocation of sample size which is optimum for all the characteristics. Cochran [24] discussed that is difficult to work out an allocation which is optimum for all characteristics unless the characteristics are highly correlated and the variation between stratum variance is very small. Compromise allocation is based on such criteria. Holmberg [25] addressed the problem of compromise allocation in multivariate stratified sampling by taking into consideration the minimization of sum of variances or coefficients of variation of population parameters and minimization of sum of efficiency loss which may result due to increase in variance because of using the compromise allocation.

The solution of a problem needs some compromise allocation criteria which make the allocation optimum for all characteristics. For example an allocation which minimizes the trace of variance-covariance matrix of the estimator of population mean or that minimizes the weighted average of variances or that maximizes the total relative efficiency of the estimators as compared to corresponding individual optimum allocation Varshney et al. [21]. Many authors Haseen et al. [26], Khan et al. [18], Kokan [27], Folks and Antle [9], Dalenius [8], Ghosh [10], Ali [23], Khan et al. [20], Ansari et al. [1], Guddat et al. [11], Haimes et al. [12], Hiller and Lieberman [13], Khan et al. [17], Charnes et al. [6], Charnes et al. [4], Khan et al. [19], Bethel [2], Chromy [7] and Khowaja [28] used different compromise criteria to solve allocation problem in stratified sampling.

This paper is organized as follows: Section 2 explains the model for the cost function. In Section 3, we formulate the problem. Section 4 explains lexicographic programming. The results and discussions are given in Section 5.

2- Quadratic Cost Functions

The cost of survey is a major factor of sample allocation to various strata. Tschuprow [30] and Neyman [29] proposed an allocation procedure that minimize variance of sample mean under a linear cost function of sample size $n = \sum_{h=1}^L n_h$ in stratified random sampling. Neyman [29] used Lagrange multiplier optimization technique to get optimum sample size for single variable under study. The linear cost function used in stratified random sampling in case of the non-response is given as;

$$C_h = C_{h0}n_h + C_{h1}n_{h1} + C_{h2}u_{h2} \quad (1)$$

Where C denotes total budget available for survey and C_{h0} be the per unit cost of selecting n_h units. C_{h1} be the per unit enumerating cost of n_{h1} units. C_{h2} be the per unit enumerating cost of u_{h2} units from the non-respondents. The expected values of $n_{h1} = W_{h1}n_h$ and $u_{h2} = W_{h2}n_h/k_h$,

where $(h = 1; 2; \dots; L)$ represent measurement per unit cost in the h^{th} stratum, and n_h is number of sample units selected from the h^{th} stratum.

Considering a quadratic cost function, including measurement unit cost and traveling cost within strata as Beard wood et al. [3] proposed the shortest route among k randomly allocated sampling units in the region is asymptotically proportional to $p\sqrt{k}$ for a large k . Varshney et al. [21] used a quadratic cost function for large sample size given in (2).

$$C = c_0 + \sum_{h=1}^L c_h + \sum_{h=1}^L \tau_h \sqrt{n_h} \quad (2)$$

where τ_h is travel cost for a unit within the h^{th} stratum.

Equation (2) can further be extended to case of presence of the non-response given in (3).

$$\sum_{h=1}^L (t_{h0} + t_{h1}W_{h1}) \sqrt{n_h} + \sum_{h=1}^L t_{h2}\sqrt{u_h} \leq C_0 \quad (3)$$

Where t_{h1} is travel cost for the respondents unit within the h^{th} stratum and t_{h2} is travel cost for the non-respondents unit within the h^{th} stratum. And u_{h2} is the sub sample from non-respondents units.

3- Optimum Allocation Techniques

The different techniques to solve the multi objective programming problem of multivariate stratified sampling in case of non-response are explained below.

3.1- Individual Optimum Technique

It is an allocation technique that optimize coefficient of variation of one characteristic of population among $Y_j(j = 1; 2; \dots; p)$ characteristics and use that allocation for estimating other characteristics of the population.

Let V_j^* be the optimum value of objective function V_j obtained by solving the following integer nonlinear mathematical programming problem (INMPP).

Minimize V_j

Subject to

$$\sum_{h=1}^L (C_{0h} + C_{h1}W_{h1})n_h + \sum_{h=1}^L C_{h2} u_h \leq c_0 \quad (4)$$

$$2 \leq n_h \leq N_h$$

$$2 \leq u_h \leq n_{h2}$$

3.2 The Goal Programming Technique

Charnes et al. [4], Charnes et al. [6], Charnes and Cooper [5] and Ijiri [14] used the goal programming technique for multi-objective optimization problems.

We can use the goal programming technique when all information about the characteristics are given and the importance of each characteristics is known. Formulation under goal programming technique can be written as:

Minimize (V_1, V_2, \dots, V_p)

Subject to

$$\sum_{h=1}^L (C_{0h} + C_{h1}W_{h1})n_h + \sum_{h=1}^L C_{h2} u_h \leq c_0 \tag{5}$$

$$2 \leq n_h \leq N_h$$

$$2 \leq u_h \leq n_{h2}$$

n_h and u_h are integers, for all h and j .

3.3- Extended Lexicographic Goal Programming Technique

Romero [31] starts reviewing the satisfying philosophy of Goal Programming (GP) and interpreting their solutions from the point of view of the utility theory. This interpretation leads to a very general optimization structure called Extended Lexicographic Goal Programming (ELGP). It is then demonstrated that there are a significant number of Multiple Criteria Decision Making (MCDM) approaches that, from a logical point of view, can be reduced to the ELGP structure.

Minimize $V_j = f(n_{jh}, u_{jh})$

Subject to

$$\sum_{h=1}^L (C_{0h} + C_{h1}W_{h1})n_h + \sum_{h=1}^L C_{h2} u_h \leq c_0 \tag{6}$$

$$2 \leq n_h \leq N_h$$

$$2 \leq u_h \leq \hat{n}_{h2}$$

$$n_{jh} \text{ and } u_{jh} \text{ are integers and } n_{jh} \in F; h = 1, 2, \dots, L$$

Note that in this generic form no assumptions have yet been made about the nature of the decision variables of goals. The decision maker(s) sets a real target level for each goal which is denoted by V_j (generally an individual optimal of the j^{th} objective). This then leads to the basic formulation of the j^{th} goal:

$$\hat{V}_j + d_j^- - d_j^+ = V_j^*$$

Where d_j^- and d_j^+ are -ve and +ve deviational variables.

The utility formulation of the Archimedean and MINMAX (Chebyshev) GP models undertaken in the preceding section suggests the following generalization:

$$\text{Minimize } (1 - \rho)D + \rho \sum_{j=1}^p [f_j(\underline{W}_{1j}d_j^-, \underline{W}_{2j}d_j^+)]$$

Subject to

$$\left[\left(\underline{W}_{1j}d_j^-, \underline{W}_{2j}d_j^+ \right) \right] \leq D \tag{7}$$

$$\hat{Z}_j + d_j^- - d_j^+ (\leq \text{ or } \geq) Z_j^*$$

$$\hat{n}_j \in F$$

$$n_{jh} \text{ and } u_{jh} \text{ are integers and } n_{jh} \in F; h = 1, 2, \dots, L$$

Where parameter W_1 and W_2 are the weights reflecting preferential and normalizing purposes attached to the negative and positive deviation variables of j^{th} goal, respectively. Parameter weights the importance attached to the minimization of the weighted sum of unwanted deviation variables and $0 \leq W \leq 1$. Here $f_j(j = 1, 2, \dots, q)$ are goals and F is the feasible space.

Integer nonlinear programming problems have a small feasible solution grid and we are already compromising on allocating sample size. This will help us to find feasible and optimal solution considering larger grid using this relaxation.

4- Application

In stratified random sampling every stratum is divided into two mutually exclusive groups of respondents and non-respondents, with N_{h1} size of respondents and $N_{h2}(N_{h2} = N - N_{h1})$ size of non-respondents in the h^{th} stratum. We select a sample of size n from the given population n_h of the respondents units and u_h from the non-respondents units are selected from N_h units in the h^{th} stratum such that $\sum_{h=1}^L n_h = n$. let $p \geq 2$ the characteristics are defined one of the population unit and the estimation of the p unknown population mean $Y_j, j = 1, 2, \dots, p$ is of interest.

Necessary formula:

$$\bar{Y}_j = \frac{1}{N} \sum_{h=1}^L \sum_{i=1}^{N_h} y_{jhi}, \bar{x}_{jh} = \frac{1}{n_h} \sum_{i=1}^{n_h} x_{jhi} \bar{X}_j = \frac{1}{N} \sum_{h=1}^L \sum_{i=1}^{N_h} x_{jhi} \bar{Y}_{jh} = \frac{1}{N_{h2}} \sum_{i=1}^{N_h} y_{jhi}$$

$$S^2_{xjh} = \frac{1}{n_h - 1} \sum_{i=1}^{n_h} (x_{jhi} - \bar{X}_{jh})^2 S^2_{yjh} = \frac{1}{n_h - 1} \sum_{i=1}^{n_h} (y_{jhi} - \bar{Y}_{jh})^2$$

$$S_{xyjh} = \frac{1}{N_h - 1} \sum_{i=1}^{n_h} (x_{jhi} - \bar{X}_{jh})(y_{jhi} - \bar{Y}_{jh})$$

Using Hansen [22] rule, as the non-response is in study variable Y_{jhi} .

Let $\bar{y}_{*jh} = (n_{h1}\bar{y}_{jh1} + n_{h2}\bar{y}_{jh2})/n_h$ be an unbiased estimator of the population mean \bar{Y}_{jh} where \bar{y}_{h1} the mean of respondents sampling units n_{h1} and \bar{y}_{jh2} is the mean of non-respondents sample units u_h .

The traditional regression estimator is $\bar{y}_{j,lrs} = \sum_{h=1}^L W_h \bar{y}_{lrs,jh}$ where $\bar{y}_{j,lrs} = \bar{y}_{jh} + b_{jh}(\bar{X}_{jh} - \bar{X}_{jh})$ and b_{jh} is sample regression coefficient.

The MSE of $\bar{y}_{j,lrs}$ is:

$$MSE(\bar{y}_{j,lrs}) = \sum_{h=1}^L W_h^2 \left(\frac{1}{n_h} - \frac{1}{N} \right) (S_{yjh}^2 - 2\beta_{jh} S_{xyjh} - 2\beta_{jh}^2 S_{xj}^2) + \sum_{h=1}^L W_h^2 \left(\frac{W_{h2}}{\mu_h} - \frac{1}{n_h} \right) W_{h2} S_{yjh}^2$$

Where

\bar{y}_{jh} = Sample Mean of j^{th} study characteristic in h^{th} stratum.

\bar{x}_{jh} = Sample Mean of j^{th} auxiliary characteristic in h^{th} stratum.

\bar{X}_{jh} = population Mean of j^{th} auxiliary characteristic in h^{th} stratum.

S^2_{yjh} = population Variance of j^{th} study characteristic in h^{th} stratum.

S^2_{xjh} = population Variance of j^{th} auxiliary characteristic in h^{th} stratum.

S^2_{yxjh} = population Covariance between the j^{th} study and the j^{th} auxiliary characteristic in h^{th} stratum, and $\beta_{jh} =$

S^2_{yxjh}/S^2_{xjh} is population regression coefficient.

Now let

$$(S_{yjh}^2 - 2\beta_{jh} S_{xyjh} - 2\beta_{jh}^2 S_{xj}^2) = \rho_{jh}$$

Then the above equation can be written as:

$$MSE(\bar{y}_{j.lrs}) = \sum_{h=1}^L \frac{W_h^2 p_{jh}}{n_h} - \sum_{h=1}^L \frac{W_h^2 p_{jh}}{N_h} + \sum_{h=1}^L W_h^2 \left(\frac{W_{h2}}{\mu_h} - \frac{1}{nh} \right) W_{h2} S_{yjh}^2$$

As the units of all characteristics measured are not same, therefore it needs to use an estimate which is free from unit measurement. So coefficient of variation is used instead of MSE.

$$CV_{(\bar{y}_{j.lrs})} = \frac{\sqrt{MSE(\bar{y}_{j.lrs})}}{\bar{Y}}$$

$$CV_{(\bar{y}_{j.lrs})} = \sqrt{\sum_{h=1}^L \frac{W_h^2 p_{jh}}{n_h} - \sum_{h=1}^L \frac{W_h^2 p_{jh}}{N_h} + \sum_{h=1}^L W_h^2 \left(\frac{W_{h2}}{\mu_h} - \frac{1}{nh} \right) W_{h2} S_{yjh}^2}$$

$$CV_{\bar{y}_{j.lrs}}^2 = \sum_{h=1}^L \frac{W_h^2 p_{jh}}{n_h} - \sum_{h=1}^L \frac{W_h^2 p_{jh}}{N_h} + \sum_{h=1}^L W_h^2 \left(\frac{W_{h2}}{\mu_h} - \frac{1}{nh} \right) W_{h2} S_{yjh}^2$$

A sample size $n = \sum_{h=1}^L n_h$ for $(h = 1, 2, \dots, L)$ is determined using proposed Quadratic cost function in Eq.3 that minimizes coefficients of variation of the estimator of population mean for each characteristics $Y_j (j = 1, 2, \dots, Q)$. This problem is formulated in multi-objective integer nonlinear programming as:

Minimize (V_1, V_2, \dots, V_p)

Subject to

$$\sum_{h=1}^L (C_{0h} + C_{h1} W_{h1}) n_h + \sum_{h=1}^L C_{h2} u_h \leq c_0$$

$$2 \leq n_h \leq N_h$$

$$2 \leq u_h \leq n_{h2}$$

n_h and u_h are integers, for all h and j .

Allocation through Individual optimum technique

Minimize V_j Subject to

$$\sum_{h=1}^L (C_{0h} + C_{h1} W_{h1}) n_h + \sum_{h=1}^L C_{h2} u_h \leq c_0$$

$$2 \leq n_h \leq N_h$$

$$2 \leq u_h \leq \hat{n}_{h2}$$

n_{jh} and u_{jh} are integers and $n_{jh}; h = 1, 2, \dots, L$

Allocation through Extended lexicographic goal programming

$$\text{Minimize } (1 - \rho)D + \rho \sum_{j=1}^p [(d_j^-, d_j^+)]$$

Subject to

$$[(d_j^-, d_j^+)] \leq D$$

$$\hat{V}_j + d_j^- - d_j^+ \leq V_j^*$$

$$\sum_{h=1}^L (C_{0h} + C_{h1} W_{h1}) n_h + \sum_{h=1}^L C_{h2} u_h \leq c_0$$

$$2 \leq n_h \leq N_h$$

$$2 \leq u_h \leq n_{h2}$$

$$n_{jh} \text{ and } u_{jh} \text{ are integers and } n_{jh}; h = 1, 2, \dots, L$$

Where $d_j(j= 1, 2, \dots, p)$ are the deviation variables.

Here V_j are the optimum values obtained from extended goal programming.

5- Numerical Illustration

The data are taken from the agricultural census 2007 conducted by National Agricultural Statistics Service, USDA, and Washington D.C. (Source: www.agcensus.usda.gov).

- Y_1 = Corn harvested in 2007.
- Y_2 = Soybean harvested in 2007.
- X_1 = Corn harvested in 2002.
- X_2 = Soybean harvested in 2002.

Where Y_1, Y_2 are study variables and X_1, X_2 are auxiliary information.

Here $\bar{Y}_1 = 22698622:75$ and $\bar{Y}_2 = 4306561:045$. It is assume that the total cost of the survey is 331 units. The last 27, 30, 27 and 20 percent values consider as non-response in each stratum respectively.

There are four strata in the population. The complete data are shown in APPENDIX.

5.1 Results and Discussion

In the three allocation techniques, extended lexicographic goal programming (ELGP) gives minimum values of CV than the other two techniques. Extended lexicographic goal programming set two additional constraints to bound Coefficients of variation maximum to their individual optimum values. Using an arbitrary weight $\theta = 0:4$ for unwanted sum of deviations from individual optimum values and $(1 - \theta) = 0:6$ for maximum deviation from utility, we minimize the goal objectives or achievements function under originally defined cost and decision variables constraints. By changing arbitrary weight θ , different results are expected (see Table 1)

Table 1: Compromise Allocations and Corresponding Values of the Objective Functions Obtained by Different Techniques

Allocation	Individual opt. technique	Goal prog.	Extended lexico.
(n_1, u_1)	(14, 3)	(17, 4)	(17, 4)
(n_2, u_2)	(35, 10)	(34, 10)	(36, 10)
(n_3, u_3)	(17, 4)	(13, 3)	(14, 3)
(n_4, u_4)	(8, 2)	(10, 2)	(10, 2)
CV_1	03100	0.03040	0.3020
CV_2	03010	0.02910	0.2900

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Appendix

Table 2: Summary Statistics of Data

h	N_h	W_h	W_{h1}	W_{h2}	$\hat{\beta}_{1h}$	$\hat{\beta}_{2h}$	S_{y1h}^2	S_{y2h}^2	S_{y1h2}^2
1	22	0.222	0.73	0.27	9833710.37	2175050.54	5.76×10^{13}	1.67×10^{12}	7.80×10^{13}
2	40	0.404	0.70	0.30	17318227.95	2308431.77	1.21×10^{14}	2.50×10^{12}	1.22×10^{14}
3	24	0.242	0.73	0.27	11778330.92	2876092.92	5.57×10^{13}	3.58×10^{12}	2.67×10^{13}
4	13	0.131	0.80	0.20	11546442.92	2980656.00	7.08×10^{13}	4.44×10^{12}	4.01×10^{13}

S_{y2h2}^2	x_{1h}	x_{2h}	S_{x1h}^2	S_{x2h}^2	S_{yx1h}	S_{yx2h}
1.48×10^{12}	0.591	0.591	0.253	0.267	2622322.77	550824.49
2.807×10^{12}	0.575	0.475	0.255	0.242	4340658.42	590425.82
3.02×10^{12}	0.50	0.542	0.259	0.167	3072607.89	745074.54
6.28×10^{11}	0.308	0.461	0.269	0.333	2664563.75	3489303.00